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Development and Validation of Decision Rules to Guide Frequency of Monitoring CD4 Cell Count in HIV-1 Infection before Starting Antiretroviral Therapy

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Abstract

Background: Although CD4 cell count monitoring is used to decide when to start antiretroviral therapy in patients with HIV-1 infection, there are no evidence-based recommendations regarding its optimal frequency. It is common practice to monitor every 3 to 6 months, often coupled with viral load monitoring. We developed rules to guide frequency of CD4 cell count monitoring in HIV infection before starting antiretroviral therapy, which we validated retrospectively in patients from the Swiss HIV Cohort Study.

Methodology/Principal Findings: We built up two prediction rules (“Snap-shot rule” for a single sample and “Track-shot rule” for multiple determinations) based on a systematic review of published longitudinal analyses of CD4 cell count trajectories. We applied the rules in 2608 untreated patients to classify their 18 061 CD4 counts as either justifiable or superfluous, according to their prior $\geq 5\%$ or $< 5\%$ chance of meeting predetermined thresholds for starting treatment. The percentage of measurements that both rules falsely deemed superfluous never exceeded 5%. Superfluous CD4 determinations represented 4%, 11%, and 39% of all actual determinations for treatment thresholds of 500, 350, and $200 \times 10^6/L$, respectively. The Track-shot rule was only marginally superior to the Snap-shot rule. Both rules lose usefulness for CD4 counts coming near to treatment threshold.

Conclusions/Significance: Frequent CD4 count monitoring of patients with CD4 counts well above the threshold for initiating therapy is unlikely to identify patients who require therapy. It appears sufficient to measure CD4 cell count 1 year after a count > 650 for a threshold of 200, > 900 for 350, or > 1150 for $500 \times 10^6/L$, respectively. When CD4 counts fall below these limits, increased monitoring frequency becomes advisable. These rules offer guidance for efficient CD4 monitoring, particularly in resource-limited settings.

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Introduction

The CD4 lymphocyte count, currently regarded as the best prognostic marker for the development of AIDS, is a major criterion to decide on initiation of antiretroviral therapy. In Western countries therapy for HIV-1 infection is recommended when the CD4 count falls below $350 \times 10^6/L$, still above the cut-off of $200 \times 10^6/L$ that strongly predicts AIDS, while evidence from non-randomized studies support treatment initiation below $500 \times 10^6/L$ [1]. Other patient characteristics may encourage earlier initiation: very high counts of circulating viral particles (over 100 000 copies/mL), rapidly falling CD4 counts (more than $100 \times 10^6/L$ per year), long-lasting inflammatory symptoms or

comorbidities [1]. Postponing treatment would otherwise be traditionally recommended [2]. The current trend is however to offer treatment to patients before their CD4 count reaches concerning levels, to prevent the deleterious effects of uncontrolled HIV-1 virus proliferation, which is possibly more hazardous than the albeit non-negligible adverse effects of antiretroviral drugs [3,4,5,6]. In countries with limited resources, a threshold CD4 count of $200 \times 10^6/L$ is commonly used, although recent guidelines and observations also suggest earlier treatment [7,8].

While a patient's CD4 count is above whatever threshold at which treatment will be started, repeated monitoring of the count is necessary. However, the *frequency* with which such monitoring should be undertaken is not currently clear, and monitoring

strategies have not been evaluated in clinical trials. Guidelines do not give explicit recommendations about monitoring frequency during the pre-treatment phase [1], although most clinicians measure the CD4 count once every 3–6 months and some also monitor viral load. Overuse of costly determinations is undesirable. However, timely introduction of antiretroviral therapy is essential in preventing AIDS and death, and underuse of monitoring in developing countries has detrimental consequences [9,10]. Two recent simulation studies concluded that improved HIV detection and pre-treatment CD4 monitoring in resource-poor settings could save several life-years per person taking antiretroviral treatment and could be cost-effective by preventing opportunistic infections [11,12,13]. Observations from the Netherlands have shown that under-screening and under-monitoring are also problematic in Western countries [14].

Our interest in evidence-based monitoring in chronic medical conditions [15] led us to analyse critically the performance of CD4 cell counts and viraemia in monitoring the pre-treatment phase of HIV infection. Our aim was to develop rational recommendations for the desirable monitoring frequency of those markers before therapy, based on published data, and to validate them in a cohort of treatment-naïve patients.

Methods

Literature review

We conducted a systematic review of the natural evolution and variability in CD4 cell count and viral load, to base our decision rules on the best evidence available from observational studies. Our aim was to summarize suitable descriptors of the longitudinal evolution of both these markers in HIV-infected patients followed up in prospective cohorts while receiving no antiretroviral treatments. We searched Medline and EMBASE for “CD4 OR viraemia OR viral load”, associated with “regression OR longitudinal OR slope OR monitoring”. We also examined the bibliographies of all relevant papers. The literature search was conducted in 2008 and updated in 2010. From among various statistical approaches to longitudinal analysis of HIV-1 biomarkers, we chose the most widely used and readily applicable, based on mixed-effects (multilevel) linear modelling. For CD4 counts, a majority of analyses used square-root transformation and thus only those ones were included; for viral load all used logarithmic transformation. The population parameters describing the evolution of biomarkers of HIV-1 were extracted, averaged, and rounded. Their prognostic value for the development of AIDS and prediction of death has been addressed in a large meta-analysis [16].

Elaboration of monitoring rules

Our literature review confirmed that a mixed-effects linear model of square-root transformed CD4 counts was most often used to describe individual trajectories in untreated patients, beyond the acute changes observed during the few months after primary infection. According to this model, the square root of the CD4 count falls along a linear mean trajectory starting from a subject-specific baseline (set point a_i) and is characterized by a subject-specific slope (decline rate b_i); actual CD4 counts depart from this line because of random fluctuations, laboratory imprecision, and model incomplete accuracy. The subject-specific baseline (a_i) and slope (b_i) are considered as random variables normally distributed around average population values (α and β), while the deviations of actual CD4 counts from the subject-specific line are considered to be independently and normally distributed around zero. The subject's specific slope and intercept represent

the “signal” hidden by the “noise” of within-subject fluctuations. As the fall in CD4 count depends on the individual slope and the time interval, measurements made close together capture only short-term variability and contain little information about the true slope. On the other hand, multiple determinations will refine evaluation of the subject's true current state, which may be advantageous in making therapeutic decisions, especially near the threshold for starting treatment. Monitoring decisions will therefore vary according to the distance of the patient from the threshold for antiretroviral treatment. The definition of such a threshold represents a peculiar aspect of CD4 monitoring [17].

We therefore designed two rules to guide decisions on CD4 monitoring frequency, detailed in Appendix S1 and illustrated in Figure 1:

“**Snap-shot rule**”. The first rule applies to a *single* CD4 measurement. The aim is to determine the time to the next observation, which has a given probability of being below the decision threshold value; in other words, to determine a time at which the likelihood of finding a clinically relevant result becomes non-negligible. If the observation at this time does not reach the decision threshold, the rule can be applied recursively and will estimate shorter and shorter intervals until the rule loses its usefulness; then either the *Track-shot rule* below or frequent CD4 monitoring become necessary. The *Snap-shot rule* is thus mainly designed for relatively high CD4 cell counts.

“**Track-shot rule**”. If *multiple* CD4 measurements have been performed in a subject, it may be worth using all those results to estimate the individual trajectory. A Bayesian approach is used to combine individual observations with published population estimates, directly inspired by a method developed for interpreting serum digoxin concentrations [18]. Once maximum likelihood estimates of the subject's specific slope and intercept have been determined, an approach very similar to the *Snap-shot rule* can be used to determine the suitable time for scheduling the next measurement, i.e. the nearest likely time when a new observation will reach the decision threshold with a given probability. As before, as CD4 counts come near to the decision threshold, this rule loses its usefulness and frequent CD4 monitoring becomes necessary.

Both rules can be applied either using global population values for the average baseline and slope, or taking into account individual characteristics that affect the CD4 trajectory, such as the viral load and age. The average population parameter values, their inter-individual variances, and the residual variance of square root CD4 count departures from a linear trajectory were drawn from the literature review.

Viral load is not commonly used as a biomarker for disease progression, due to prominent variability and absence of clear time trend over the asymptomatic phase of HIV-1 infection in a fair number of patients. Moreover, viral load is not usually interpreted in terms of a decision threshold for starting antiretroviral therapy. This leaves little role for specific rules to guide determination frequency, as confirmed by our attempt to apply a similar approach to this marker (Appendix S2 and Figure S1).

Validation study

To test their performance with actual clinical data, we applied both decision rules to an unselected series of HIV-1 infected

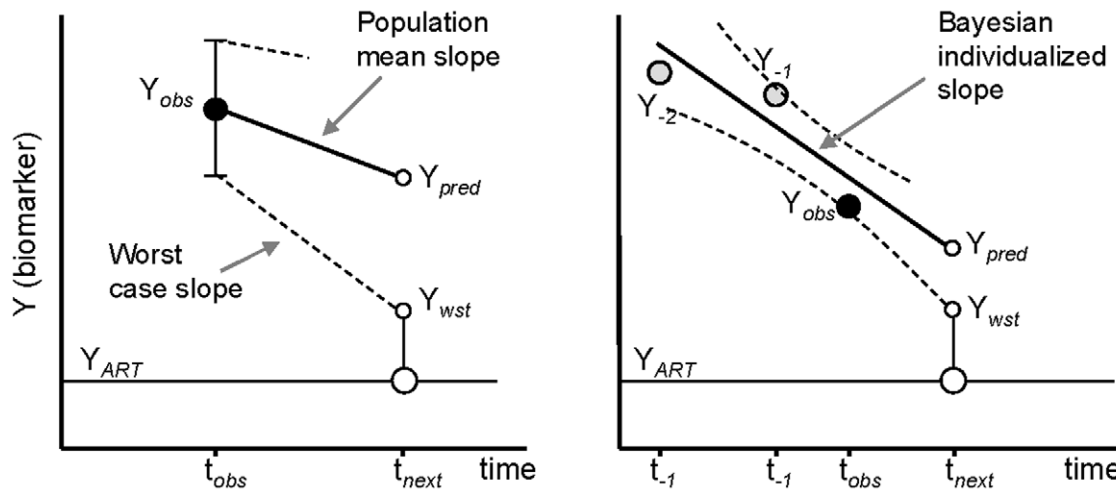


Figure 1. Illustration of CD4 monitoring decision rules. The *Snap-shot rule* (left) uses a single observation Y_{obs} (square root transformed) of the biomarker at time t_{obs} . The *Track-shot Rule* (right) uses Y_{obs} plus one or more previous observations Y_{-1} , Y_{-2} etc. available at times t_{-1} , t_{-2} etc. The suitable time for next measurement, t_{next} , is when the predicted value has some minimal probability P to reach the decision limit for antiretroviral therapy Y_{ART} . Appropriate standard normal deviates z_p are used to weight the within-subject and between-subject dispersions, σ_e and σ_b respectively, according to the level chosen for P . The relevant prediction therefore depends on a worst-case scenario (Y_{wst}) rather than average population prediction (Y_{pred}). This illustration is schematic, as it is variances that are actually summed, not standard deviations. doi:10.1371/journal.pone.0018578.g001

patients. The multi-centre Swiss HIV Cohort Study (SHCS), established in 1988, includes about a half of all HIV-positive individuals in Switzerland (www.SHCS.ch). Detailed clinical data are obtained from the subjects on recruitment and then at visits scheduled at least 6-monthly; all CD4 and viral load determinations are entered into the database. All patients give their written consent to epidemiological analysis of their follow up data. This validation study was approved by the SHCS scientific board.

Of the 9570 subjects recruited during the past 12 years (1 January 1996 to 31 May 2008), 5551 have undergone CD4 and viral load monitoring before starting antiretroviral therapy. The data from patients with at least *two* such recorded values were used to validate the rules outlined above, applied recursively from the first value. To prevent selection bias in favour of slow progressors, cases were included only if their first biomarker value was recorded after 1996. We did not adjust for right truncation due to treatment initiation, interruption in a subject's follow-up beyond 2008, or other causes. For all included cases, we used the rules to determine an optimal length of time between measures, based on one or more initial measures and predefined treatment decision thresholds. We started from the first CD4 count recorded, with P set to 0.05 (i.e. re-testing was deemed *superfluous* before having at least a 5% chance to reach the treatment initiation threshold). We also tested both rules with P set to 0.1. Subsequent CD4 determinations performed before the time indicated by the rule (t_{next}) were counted, compared with the threshold, and discarded as superfluous. The *Snap-shot rule* was then reiterated on the next justifiable (i.e. non-superfluous) measurement, done after t_{next} and so on. The *Track-shot rule* took into account only previous CD4 values not discarded. Treatment thresholds were set at CD4 counts of 200, 350, and $500 \times 10^6/L$.

The main objective of the study was to verify the adequacy of the classification of CD4 measures by both rules. Concretely, we aimed to check that not more than 5% of the CD4 results deemed superfluous would actually reach the threshold, using a predefined P level of 0.05; or not more than 10% using a P of 0.1. Calculations were performed using Excel (Microsoft, Redmond WA, 2003) and STATA software (v. 10, StataCorp, College-Station TX, 2007).

Results

Literature review

Most published descriptions of CD4 cell counts and viral load evolution in untreated HIV infection come from cohort data. We screened 149 abstracts addressing the topic of the rate of fall in CD4 counts in untreated HIV-1 infection, and we identified 40 publications that described mathematically the natural evolution of CD4 counts in untreated HIV-1 infected patients. Among them, 11 did not provide usable parameters, while 19 used longitudinal models, data transformations, or parameterizations that were not relevant to our approach (see PRISMA Flow Diagram S1). Thus, we included 8 analyses that provided suitable estimates of parameters describing the fall in CD4 count in the square-root scale according to a mixed-effects linear model [19,20,21,22, 23,24,25,26] (Table 1); we also included the summary estimates from two similar reviews [11,27]. CD4 count has high variability between measures, even taken a few hours or days apart, with coefficients of variation of 13–26% [28,29]. Square-root transformation simultaneously renders average trajectories approximately linear and residual errors approximately Gaussian, and is the method that has been used most often with CD4 count data [20,30]. Patients are reported to differ regarding both their baseline (set point) and slope (rate of fall) of CD4 count; non-progressors remain stable over years [31] while rapid progressors lose large numbers of cells over short periods [32]. Slopes tend to be steeper in individuals who start from a high baseline, and are mainly correlated with viral load [32,33,34,35] and increasing age [19,36]; they are also affected by HIV strain [37], transmission route [38], gender [35], race [35], pregnancy [39], genetics [40], and immune reactivity [41]. However, all those factors explain only a small percentage of the overall variability in the rate of fall in CD4 count (less than 10% for viral load and age) [42]. Average CD4 decline rates have been reported to be similar between African and Western countries [43], although recent observations suggest slightly slower rates in non-white individuals [44].

There was a moderate degree of heterogeneity across the estimates shown in Table 1. To elaborate our general monitoring

Table 1. Published estimates of CD4 cell count decay rate and variability in untreated HIV-infected individuals.

| Reference | β | σ_b | α | σ_a | σ_e | ρ_{ab} | N-n | duration |
|----------------------|-------------------------|-------------------|-----------------------|------------|---------------------|-------------|------------------|----------|
| DeGruttola 1991 [20] | 2.1 | 1.1 | 33 | 4 | 3 | -0.9 | 495-5 | 3 |
| Lange 1992 [23] | 1.6 | 0.6 | 30 | 2 | - | -0.5 | 327-8 | 3.5 |
| Faucett 1996 [21] | 2.3 | 2.4 | 25 | 10 | 2.7 | -0.6 | 109-6 | 4.1 |
| Lepri 1997 [22] | 1.7 | 2.8 ^{oo} | 26 | 6 | - | - | 1021 | >10 |
| Touloumi 1998 [19] | 1.4 | 1.2 | 25 | 5 | - | - | 137-9 | 10 |
| Laurent 2002 [24] | 1.3 | - | 26 | - | - | - | 331-3 | 2 |
| CASCADE 2003 [25] | 1.3 to 1.7 [*] | 1.4 | 23 to 29 ^o | 6 | 3 | -0.4 | 5739-9 | 5.2 |
| Taffé 2008 [26] | 2.1 | 1.2 | 30 | 3 | 6 | -0.4 | 4217-6 | 3 |
| Phillips 2007 [27] | 0 to >2 [†] | 0.8 | 39 | 2 | - | - | (review) | |
| Hallett 2008 [11] | 1.3 to 2 [‡] | 1 | 26 | 1 | (50 ^{**}) | - | (review) | |
| Average | 1.8 | 1.2 | 30 | 5 | 4 | -0.5 | (rounded values) | |

Parameters derived from linear mixed effect model in the square root scale. β : slope and σ_b : inter-individual slope variability in $(10^6/L)^{0.5}/\text{year}$; α : intercept, σ_a : inter-individual intercept variability and σ_e : intra-individual variability in $(10^6/L)^{0.5}$; ρ_{ab} : correlation between α and β ; N: number of individuals, n: average number of samples per individual; average duration of observation in years.

*slope according to age (15–20: 1.30, 20–30: 1.53, 30–40: 1.73, >40: 1.67) and to symptoms at pre-infection (present: +0.26).

^ointercept depending on subgroup (sex, age, intravenous drug use or haemophilia).

[†]slope depending on viral load (0 at $<10^3$, then 0.016, 0.04, 0.12, 0.4, 0.8 and 1.6 for every $10^{0.5}$ step up to 2.0 at $>10^6/\text{mL}$), age (+0.007/year), X4-virus shift (present: +0.25).

[‡]slope depending on age (<35: 1.3, >35: 2).

**standard deviation of a uniform distribution in the untransformed CD4 count scale.

^{oo}outlier value discarded from average calculation.

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rules, we simply took the rounded average values. We also performed sensitivity analyses to check the robustness of the rules towards alteration in the parameters.

CD4 monitoring recommendations

The operation of the *Snap-shot rule* to determine the suitable time for CD4 remeasurement is illustrated in Figure 2 using the “variogram” approach [45] based on equation 3 above. It shows the lowest CD4 count that a subsequent measurement is expected to reach with a preset probability, P , for a given result. The essence of this rule can be represented in a nomogram (Figure 3). The *Snap-shot rule* justifies delayed CD4 measurement for higher values observed:

- With a treatment threshold count of $200 \times 10^6/\text{L}$, under 5% of individuals are expected to reach the threshold within 6 months after a count of $600 \times 10^6/\text{L}$, 1 year after $650 \times 10^6/\text{L}$, 1.5 years after $700 \times 10^6/\text{L}$, 2 years after $770 \times 10^6/\text{L}$, 3 years after $950 \times 10^6/\text{L}$, or 4 years after $1100 \times 10^6/\text{L}$.
- A threshold of $350 \times 10^6/\text{L}$ has less than a 5% chance of being reached 6 months after a value of $840 \times 10^6/\text{L}$, 1 year after $900 \times 10^6/\text{L}$, 1.5 years after $970 \times 10^6/\text{L}$, 2 years after $1050 \times 10^6/\text{L}$, or 3 years after $1250 \times 10^6/\text{L}$.
- A threshold of $500 \times 10^6/\text{L}$ has less than a 5% chance of being reached 6 months after a value of $1060 \times 10^6/\text{L}$, 1 year after $1140 \times 10^6/\text{L}$, 1.5 years after $1220 \times 10^6/\text{L}$, or 2 years after $1300 \times 10^6/\text{L}$.

The rule loses its usefulness in the presence of low to intermediate CD4 counts (below 550, 785, and $1000 \times 10^6/\text{L}$ for the three respective decision thresholds at $P=0.05$, and below 460, 675, and $880 \times 10^6/\text{L}$ at $P=0.1$).

The operation of the *Track-shot rule* cannot be summarized in a chart, since it is based on serial measurements. This rule individualizes to a reasonable extent the mean expectation of the

CD4 trajectory with regard to previous results. It also reduces the width of the prediction interval when many results have been recorded. But even for a trajectory defined with high precision by

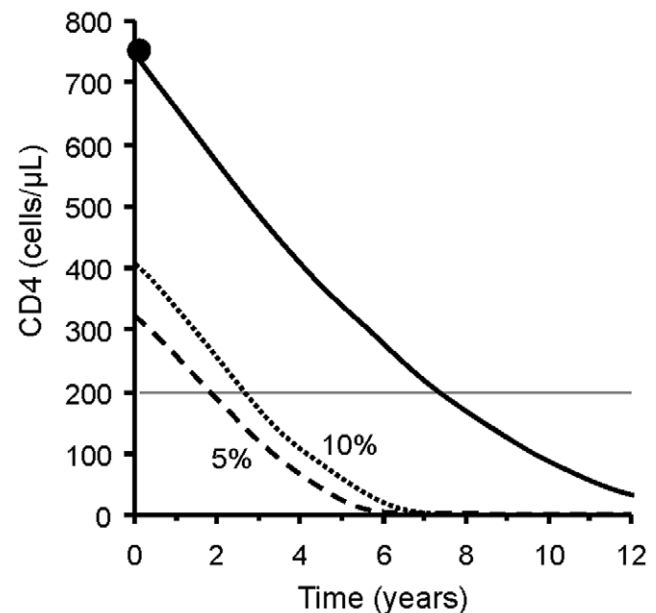


Figure 2. Variogram for the *Snap-shot rule*. The *Snap-shot rule* is applied to an initial CD4 cell count of $750 \times 10^6/\text{L}$, observed at time=0. It shows the lowest value that a subsequent measurement can be expected to reach with a probability of 5% (dashed line) or 10% (dotted line). The continuous line indicates the CD4 trajectory predicted for an average patient. The average curve will take 7.3 years to reach the $200 \times 10^6/\text{L}$ threshold (horizontal line); however, biological variability makes this outcome possible within 2 years for one patient in 20, and at 3 years for one patient in 10.

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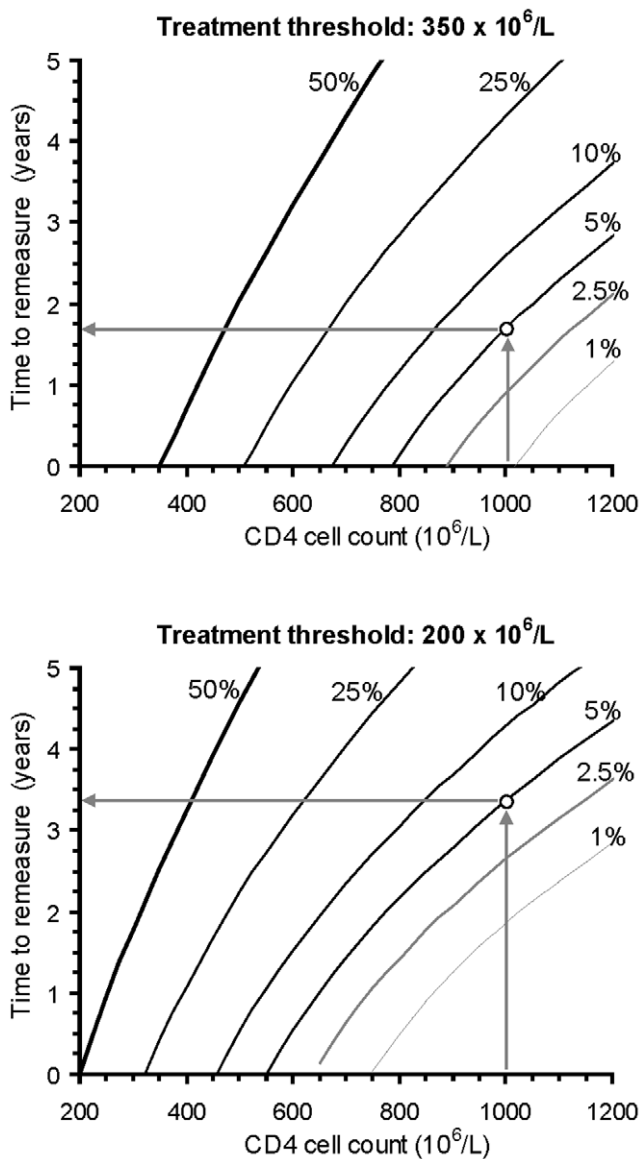


Figure 3. Nomograms for the *Snap-shot* rule. These nomograms show the time to wait before the next CD4 count determination as a function of the actual observation, at two decision thresholds to start antiretroviral therapy, with varying probabilities of observing a value at this threshold. The 50% lines correspond to average population predictions. The arrows illustrate the rule applied to an initial count of $1000 \times 10^6/L$, giving about 1.7 years to reach a count of $350 \times 10^6/L$ with a 5% chance, and about 3.4 years to reach a count of $200 \times 10^6/L$. doi:10.1371/journal.pone.0018578.g003

numerous points, this rule loses its usefulness once intra-individual variability makes it possible for the next point to reach the threshold (respectively below 430, 640, and 840 for the three thresholds at $P=0.05$, and 370, 570, and 755 at $P=0.1$). An Excel spreadsheet implementing both the *Snap-shot* rule and the *Track-shot* rule is available as supporting information (see Excel Tool S1).

Validation of the CD4 monitoring rules

In total, 2608 patients from the Swiss HIV Cohort Study were entered into the validation study (Table 2). They had 18 061 CD4 measurements over a median (interquartile range, IQR) follow-up of 559 (182–1182) days, consecutive measurements being

separated by 105 (85–171) days. The CD4 cell counts spread around a median of $427 (310–593) \times 10^6/L$, while their square-root transforms followed a fairly symmetric, bell-shaped distribution, with a mean (SD) of $20 (6.2) (\text{cell count})^{0.5}$. Two thirds of the patients started antiretroviral therapy before May 2008.

Compared with the study patients, those who could not be included because they lacked repeated CD4 testing were slightly more often women (66%, $P=0.06$), older (37, IQR 32–42 years, $P<0.001$), and intravenous drug users (29% vs 18%, $P<0.001$). Above all, they had significantly lower initial CD4 counts, with a median of 250 (111–430, $P<0.001$), and higher initial viral loads, with a median of 51 800 (10 400–175 000, $P<0.001$). Among them, 2689 (91%) started antiretroviral therapy soon after their single off-treatment CD4 determination, which presumably often represented a main criterion for treatment decision.

Application of the monitoring rules, illustrated in Figure 4, shows that the frequency of CD4 determinations can be kept low while high counts are observed. In most situations, the *Snap-shot* rule and the *Track-shot* rule give similar results.

We applied both rules iteratively to CD4 counts in the patients in the validation study. Table 3 summarizes the evaluation of the 15 453 non-initial CD4 values according to the rules, each run at two levels of probability ($P=0.05$ and 0.1) and at three treatment decision thresholds ($200, 350, \text{ and } 500 \times 10^6/L$). The *Snap-shot* rule declared fewer measurements superfluous than the *Track-shot* rule for a threshold of $200 \times 10^6/L$, an equal amount for $350 \times 10^6/L$, and more for $500 \times 10^6/L$. However, the corresponding absolute counts fell dramatically on increasing the threshold, as expected. The analysis confirmed that the percentage of CD4 results that were actually below the threshold value among determinations deemed superfluous was always lower than the preset probability level P used to run the rules, moderately so for the 350 and $500 \times 10^6/L$ thresholds but markedly for the $200 \times 10^6/L$ threshold. There were 4905 superfluous tests in 1024 patients, taken 322 (IQR 139–605) days before the time indicated by the *Snap-shot* rule run at $P=0.05$ for the $200 \times 10^6/L$ threshold; for the $350 \times 10^6/L$ threshold, there were 1724 superfluous tests in 421 patients, requested 287 (122–555) days in advance; and for the $500 \text{ cells}/\mu\text{L}$ threshold, 625 superfluous tests in 192 patients, requested 256 (102–482) days in advance. When we restricted the analysis to the CD4 results obtained within ± 3 months of the date indicated as suitable for re-measurement by the rules, the absolute numbers of values became smaller (219–2393), while the percentages of falsely superfluous CD4 results remained below the preset probability (0.68%, 2.93%, and 4.38% for the *Snap-shot* rule at $P=0.05$ at the 200, 350, and $500 \times 10^6/L$ thresholds respectively). Very similar results emerged for the *Track-shot* rule (not detailed).

There was no advantage in using the rules with individualized slopes modified according to age, initial viral load, or HIV infection route (not detailed). The *Track-shot* rule evaluated average (SD) posterior slope estimates of $1.67 (0.51) (\text{cell count})^{0.5}/\text{year}$, and intercept estimates of $26 (4.3) (\text{cell count})^{0.5}$. Its use with a non-informative prior intercept gave very similar results; conversely, a non-informative prior slope severely compromised the validity (not detailed).

Sensitivity analyses

Both rules were robust towards changes in parameter values, which did not markedly affect the outcomes. For the *Snap-shot* rule, a 25% reduction in the population average slope (β) translated into a few more measurements being declared superfluous (35% instead of 32% for a threshold of 250 cells/ μL , 4.6% instead of 4.1% for 500 cells/ μL); conversely, increasing β by 25% slightly reduced this percentage (29% instead of 32%, and 3.6% instead of

Table 2. Description of the subpopulation of patients drawn from the Swiss HIV Cohort Study for the validation study.

| Characteristic | |
|--|----------------------|
| Number of patients | 2608 |
| Men, n (%) | 1786 (68.5%) |
| Age, median (IQR) | 35 (29–41) |
| Likely source of HIV infection | |
| Heterosexual contact | 1105 (42.4%) |
| Homosexual contact | 930 (35.7%) |
| Intravenous drug use | 473 (18.1%) |
| Other | 42 (1.6%) |
| Unknown | 58 (2.2%) |
| Days since first HIV positive test, median (IQR) | 40 (14–453) |
| Initial CD4 count, median $\times 10^6/L$ (IQR) | 426 (286–605) |
| Initial viral load, median copies/mL (IQR) | 24 000 (5100–93 800) |
| Number of CD4 determinations before antiretroviral therapy, median (IQR) | 5 (3–9) |
| Initiation of antiretroviral therapy, n (%) | 1759 (67%) |
| Days to initiation of antiretroviral therapy, median (IQR) | 344 (87–889) |
| Last CD4 count before antiretroviral therapy, median $\times 10^6/L$ (IQR) | 239 (166–328) |

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4.1%); these manipulations minimally affected the rate of falsely superfluous results (all absolute changes $<0.8\%$). Increasing or decreasing the slope dispersion (σ_b) by 25% had little effect on the

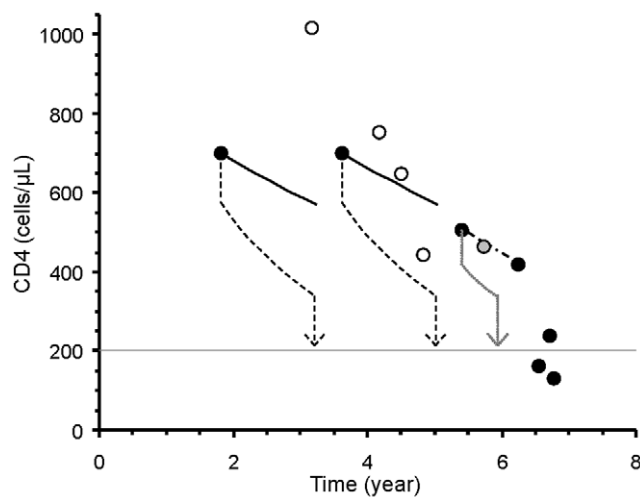


Figure 4. CD4 cell count data from an illustrative patient. The *Snap-shot rule* is applied to the first determination (closed circle) done 1.8 year after positive serology (discovered at time = 0), and indicates to repeat the test at 3.2 years (dashed arrow), when the CD4 count has 5% chance to reach 200 cells/ μL . The second value actually measured at 3 years is thus declared superfluous and discarded (open circle). The rule is reapplied to the third value actually measured at 3.6 years, and designates the 3 next determinations as superfluous. The last 6 values fall below the threshold to question necessity. The treatment is actually initiated 7 years after HIV detection, once CD4 counts reach the threshold of 200 cells/ μL . The lines indicate the average population slope. The *Track-shot rule* would give fairly similar results, except that it would indicate the 8th measurement as superfluous (dotted arrow and grey circle).

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number of either superfluous measurements (absolute changes $<1.3\%$), or falsely superfluous results (absolute changes $<0.4\%$). Modifying intra-patient variability (σ_e) had more dramatic effects: a 25% reduction in σ_e increased both the number of superfluous measurements (47% instead of 32% for a threshold of 250 cells/ μL , 8.1% instead of 4.1% for 500 cells/ μL) and the rate of falsely superfluous results (0.9% instead of 0.5% for a threshold of 250 cells/ μL , 5.7% instead of 4.8% for 500 cells/ μL), while a 25% increase in σ_e had the opposite effect. Despite this, none of the actual percentages of falsely superfluous rates, calculated with a preset P value of 0.05, exceeded 5.7%. The *Track-shot rule* was not more sensitive to changes in the choice of parameter values (not detailed).

Discussion

Monitoring CD4 cell counts in asymptomatic HIV-1 infected patients to decide when to start antiretroviral therapy is unanimously recommended and is cost-effective [11,12]. However, no recommendations have yet been formulated regarding the optimal frequency of CD4 monitoring, which most UK practitioners perform every 3–4 months [46]. Based on a review of published observations describing average CD4 decline rates and variability in populations of untreated HIV-1 infected patients, we developed two decision rules aimed at guiding CD4 monitoring decisions [17]. The *Snap-shot rule* applies to a single measurement. The *Track-shot rule* takes into account a series of CD4 results in a given patient. We validated both rules using CD4 counts collected before the start of therapy in a large cohort of HIV-1 infected patients. The rules proved reliable and robust and can therefore be used to guide the management of asymptomatic patients.

We used a similar approach to evaluate the performance of repeated viral load determination based on a systematic literature review (see Appendix S2). It confirmed that this measurement has a limited role in monitoring patients who are not taking antiretroviral therapy, as others have found [12,47].

Various approaches have been used to describe the average trend of CD4 or viral load and their inter- and intra-patient variability [48]. Since the early 1990s, mixed-effects modelling has progressively prevailed [20,30]. Both of our rules use this approach, which minimizes contamination of the parameters that describe individual trajectories by intra-individual fluctuations and measurement errors. It appropriately shrinks the estimation of slope variability, thus clearly improving predictive performance [30].

We initially expected that the *Track-shot rule* would outperform the *Snap-shot rule*, since it adduces more information. However, except for patients starting from high CD4 levels and targeting a low treatment threshold ($200 \times 10^6/L$), the *Track-shot rule* was not superior (Table 3). Relaxing prior assumptions regarding the intercept of CD4 trajectory (α) changed nothing (though it is theoretically appropriate to do so, owing to uncertainty in the HIV-1 seroconversion date in many patients). We attribute this important finding to the fact that the fall in CD4 count with time is largely dominated by large intra-individual variability (σ_e), which combines biological fluctuation, laboratory imprecision, and model inaccuracy; inter-individual slope variability plays the second role. This is further reflected by the results of our sensitivity analysis.

The observed rates of false superfluous CD4 tests were below the preset probability level (P), in particular for the low target threshold of $200 \times 10^6/L$ (Table 3). This is partly due to the more frequent repetition of CD4 testing in patients with slower rates of fall in CD4 count, while the most rapidly progressing patients were

Table 3. Validation of the two rules on all non-initial CD4 cell counts.

| Rule: | | | <i>Snap-shot rule</i> | | <i>Track-shot rule</i> | | No rule |
|---------------------------------------|--|--------------------|-----------------------|--------------|------------------------|--------------|---------|
| Probability level: | | | <i>P=0.05</i> | <i>P=0.1</i> | <i>P=0.05</i> | <i>P=0.1</i> | |
| Threshold for antiretroviral therapy: | Evaluation of informativity: | Actual CD4 result: | | | | | |
| | | ≤200 cells/μL | | | | | |
| | justifiable | ≤200 | 1154 | 1112 | 1144 | 1069 | 1179 |
| | | >200 | 9394 | 7126 | 8268 | 5313 | 14 274 |
| | superfluous | ≤200 | 25 | 67 | 35 | 110 | – |
| | | >200 | 4880 | 7148 | 6006 | 8961 | – |
| | <i>percent superfluous*</i> | | 32% | 47% | 39% | 59% | – |
| | <i>percent ≤200 among them[°]</i> | | 0.5% | 0.9% | 0.6% | 1.2% | 7.6% |
| ≤350 cells/μL | justifiable | ≤350 | 5153 | 5108 | 5152 | 5118 | 5196 |
| | | >350 | 8576 | 7416 | 8653 | 7358 | 10 257 |
| | superfluous | ≤350 | 43 | 88 | 44 | 78 | – |
| | | >350 | 1681 | 2841 | 1604 | 2899 | – |
| | <i>percent superfluous*</i> | | 11% | 19% | 11% | 19% | – |
| | <i>percent ≤350 among them[°]</i> | | 2.5% | 3.0% | 2.7% | 2.6% | 33% |
| ≤500 cells/μL | justifiable | ≤500 | 9701 | 9661 | 9717 | 9682 | 9731 |
| | | >500 | 5127 | 4608 | 5323 | 4892 | 5722 |
| | superfluous | ≤500 | 30 | 70 | 14 | 49 | – |
| | | >500 | 595 | 1114 | 399 | 830 | – |
| | <i>percent superfluous*</i> | | 4.1% | 7.7% | 2.7% | 5.7% | – |
| | <i>percent ≤500 among them[°]</i> | | 4.8% | 5.9% | 3.4% | 5.6% | 63% |

The 15 453 non-initial CD4 tests in 2608 treatment naïve patients were first classified as either justifiable or superfluous based on the rules, and then compared with the preset threshold to assess the percentage actually measured below it.

*All 95% confidence band widths <0.8% (normal approximation for binomial proportions).

[°]All 95% confidence band widths <1.8% (normal approximation for binomial proportions).

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soon to receive treatment. Besides this right truncation, our validation population was also affected by selection bias. Indeed, a significant number of patients with a more active form of HIV-1 infection never underwent off-treatment CD4 monitoring, but immediately started antiretroviral therapy. This is reflected in the lower initial CD4 counts and higher viral loads in non-included patients. This is also in accordance with the fact that the post-hoc slope estimates given by the *Track-shot rule* were distributed around a lower average than we found in our literature review. Such selection bias actually represents a reassuring argument for implementing the rules in a clinical setting, where similar selection bias is likely to occur: our rules, based on average CD4 kinetic parameters, are thus expected to perform with an extra safety margin when they are used in patients whose clinical condition allows treatment to be deferred.

The rules did not perform better when age, initial viral load or HIV infection route, which are thought to modulate the average rate of fall in CD4 count, were incorporated. This is probably because, despite their statistical significance in cohort studies, age and viral load explain only a tiny part of overall CD4 variability [34,42].

The consequences of CD4 measurement variability on therapeutic decisions were recognized in 1992 by Hoover, who suggested starting therapy after measuring two consecutive CD4 counts below the decision threshold, instead of one [28]. The phenomenon of regression to the mean, resulting from oscillation of values around their long-term trajectory, may be viewed as an argument for frequent monitoring. However, as we have shown, analysis of biological variability results in recommendations that

spare measurement resources when frequent monitoring is unnecessary, in this case after a high CD4 cell count. CD4 monitoring represents only one aspect of the management of HIV-infected patients who do not require therapy, whose follow-up frequency should be arranged according to many other factors (e.g. co-morbidities, prevention of contamination, need for psychological support).

We found that roughly 11% of all CD4 cell counts performed in untreated patients were of questionable clinical usefulness, as they had less than a 5% chance of being under $350 \times 10^6/L$. This represented 1724 measurements, or a global cost of 93 000 CHF (€58 000, £47 000, \$82 000) for 421 patients. The percentage would increase to almost 40% with the *Track-shot rule* at a treatment threshold of $200 \times 10^6/L$, which was commonly used in developing countries until very recently. Although there is no indication that our rules would perform differently in resource-poor settings, they should be validated in populations differing from Swiss HIV patients.

Finally, the main limitation of our rules is their inability to provide guidance for following CD4 cell counts when they start to approach the threshold for starting therapy. This is a consequence of the fact that the variability in CD4 count is mainly governed by the intra-individual component (σ_c). The *Snap-shot rule* ceases to be applicable once the variogram meets the treatment threshold (Figure 2). If there are multiple CD4 determinations, the *Track-shot rule* can be used later on, but loses its usefulness once intra-individual variability makes it likely that an immediately retested CD4 count will reach the treatment threshold. Beyond this point, traditional testing at about 3 month intervals is justified. A single

CD4 count below the threshold should be confirmed before starting antiretroviral therapy. In many cases will the repeat measure exceed the threshold, allowing further delay before initiation [28].

In conclusion, both theoretical arguments and observational evidence suggest that CD4 cell count monitoring, as performed in HIV-1 infected patients who do not require treatment, is currently too frequent in those whose CD4 counts are well above the treatment threshold. Infrequent measurement can safely be recommended in this subpopulation.

Supporting Information

Figure S1 Variogram for viral load monitoring. This variogram, based on the *Snap-shot rule* for an initial determination of a viral load of 1000 copies/mL (3 log units/mL), shows the highest load that a subsequent measurement can be expected to reach, with a probability of 5% (dashed line) or 10% (dotted line). The continuous line indicates the viral load trajectory predicted in an average patient, taking about 12 years to increase by 1 log unit. After 6.7 years one patient in 20, and after 9.2 years one patient in 10, can be expected to have a 2 log unit increase (i.e. to 100 000 copies/mL, arrows).

(TIF)

Appendix S1 Construction of the monitoring rules: statistical assumptions and mathematical derivations underlying the *snap-shot rule* and the *track-shot rule*.

(PDF)

Appendix S2 Evaluation of viral load monitoring: literature review and application of the rules to viral load.

(PDF)

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